INTCAL04 TERRESTRIAL RADIOCARBON AGE CALIBRATION, 0–26 CAL KYR BP

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ABSTRACT. A new calibration curve for the conversion of radiocarbon ages to calibrated (cal) ages has been constructed and internationally ratified to replace IntCal98, which extended from 0–24 cal kyr BP (Before Present, 0 cal BP = AD 1950). The new calibration data set for terrestrial samples extends from 0–26 cal kyr BP, but with much higher resolution beyond 11.4 cal kyr BP than IntCal98. Dendrochronologically-dated tree-ring samples cover the period from 0–12.4 cal kyr BP. Beyond the end of the tree rings, data from marine records (corals and foraminifera) are converted to the atmospheric equivalent with a site-specific marine reservoir correction to provide terrestrial calibration from 12.4–26.0 cal kyr BP. A substantial enhancement relative to IntCal98 is the introduction of a coherent statistical approach based on a random walk model, which takes into account the uncertainty in both the calendar age and the ¹⁴C age to calculate the underlying calibration curve (Buck and Blackwell, this issue). The tree-ring data sets, sources of uncertainty, and regional offsets are discussed here. The marine data sets and calibration curve for marine samples from the surface mixed layer (Marine04) are discussed in brief, but details are presented in Hughen et al. (this issue a). We do not make a recommendation for calibration beyond 26 cal kyr BP at this time; however, potential calibration data sets are compared in another paper (van der Plicht et al., this issue).

INTRODUCTION

As archaeologists and geoscientists strive for higher resolution and greater chronological controls in their studies, there is a need for continued refinement of the radiocarbon calibration data set to convert ¹⁴C ages into calibrated ages. This latest effort to provide a consensus calibration data set for terrestrial and marine samples is the product of the IntCal working group, which met at Queen’s University Belfast in April 2002 and at Woods Hole Oceanographic Institute in May 2003. The
IntCal working group established criteria for acceptance of data into the IntCal calibration data set including general limitations on analytical errors and acceptable scatter and specific record-dependent criteria (Reimer et al. 2002). Specific criteria for tree-ring data sets are reviewed in brief in the following section, while those for marine data sets are discussed in Hughen et al. (this issue a). The data sets were presented and discussed at the 18th International Radiocarbon Conference in Wellington, New Zealand and ratified at the Business Meeting. Suggestions from conference participants have been incorporated into the final product.

This paper focuses on the IntCal04 calibration data set for $^{14}$C ages of Northern Hemisphere terrestrial samples. However, because $^{14}$C measurements of foraminifera from the Cariaco Basin varved sediments and U-series-dated coral are the basis for the terrestrial calibration data set beyond the beginning of the tree rings at 12.4 kyr, we will discuss marine data in brief. The terrestrial calibration for the Southern Hemisphere is discussed in McCormac et al. (this issue). For IntCal04, tree-ring measurements are utilized from 0–12.4 cal kyr BP and marine data, with site-specific, regional reservoir corrections, from 12.4–26 cal kyr BP as illustrated in Figure 1. These data are combined statistically using a random walk model (RWM) described in detail in Buck and Blackwell (this issue). The calibration curve for surface mixed-layer marine samples, Marine04, is constructed from coral and foraminifera $^{14}$C measurements from 10.5–26 cal kyr BP. These data are normalized to the “global” ocean mixed layer by subtracting the difference between the regional reservoir age and the marine model ($\Delta R$) as defined by Stuiver et al. (1986) and combined via the RWM. Because marine data are sparse from 0 to 10.5 cal kyr BP, a global ocean-atmosphere box diffusion model (Stuiver and Braziunas 1993) is used to deconvolve the $^{14}$C production rate from the tree-ring data and calculate the “global” ocean mixed-layer $^{14}$C ages for this period. The construction of the calibration data set for marine samples (Marine04) is presented in full in a companion article (Hughen et al., this issue a).

![Figure 1 Schematic diagram of IntCal04 and Marine04 calibration data set construction. Tree-ring data extend from 0 to 12.4 cal kyr BP. Beyond the end of the tree rings, coral and foraminifera data are converted to the atmospheric equivalent by subtracting a site-specific reservoir correction $\Delta R$. These data are input into the random walk model (RWM) to produce IntCal04. Marine data from 10.5–26.0 cal kyr BP are normalized to the “global” ocean by subtracting $\Delta R$, the regional difference from the model ocean reservoir age of 405 yr. The “global” marine data and the output from the box diffusion model (see main text) are used with the RWM to provide our estimate of the global marine calibration curve Marine04.](image-url)
We do not make a recommendation for calibration beyond 26 cal kyr BP because disparities between 
$^{14}$C data sets are large and have not been resolved. The strengths and weaknesses of various 
$^{14}$C data sets beyond 26 cal kyr BP are discussed by van der Plicht et al. (this issue). These include 
U/Th-dated speleothems (Beck et al. 2000; Weyhenmeyer et al. 2003), macrofossils from varved 
lake sequences (Kitagawa et al. 2000), U/Th-dated aragonite from Lake Lisan (Schramm et al. 2000; 
van der Borg et al. 2004), foraminifera from marine sediments with proxies matched to the ice-core 
time scales (Bard et al., this issue; Hughen et al. 2004; Voelker et al. 2000; van Kreveld et al. 2000), 
and U/Th-dated corals (Bard et al. 1998; Cutler et al., this issue). We actively discourage the use of 
subjectively assembled “calibration” curves from disparate data sets; these are not recognized as 
“calibration” by the international community and should not be published as such.

THE INTCAL04 DATA SETS

The data sets used in the IntCal04 calibration are given in full as supplemental material on the 
Radiocarbon Web site (www.radiocarbon.org) and are also available at www.calib.org. Uncertainties 
are given for the data set ages and the calibrated or cal time scales in order that they may be com-

bined properly, since in some cases the cal time scale errors are not independent. Replicate 
$^{14}$C measurements within a laboratory or made by two or more laboratories are given separately, when 
available. These data are not necessarily completely independent estimates of the underlying cali-

bration curve, since they are derived from the same samples, but have been included for complete-

ness.

Tree-Ring Data Sets (0–12.4 cal kyr BP)

For inclusion in the calibration data set, dendrochronological dating and cross-checking of tree rings 
is required. A few exceptions were made for post-AD 1320 Douglas firs ($Pseudotsuga menziesii$), 
which have robust ring production in the Pacific Northwest environment. Calendar ages for these 
trees, which had well-known felling dates, were determined by ring counting after inspection of the 
well-developed rings by H C Fritts (Stuiver 1982). X-ray densitometry was later used to confirm the 
ring counts.

The Holocene part of the $^{14}$C calibration is based on several millennia-long tree-ring chronologies, 
providing an annual, absolute time frame within the possible error of the dendrochronology, which 
was rigorously tested by internal replication of many overlapping sections. Whenever possible, they 
were cross-checked with independently established chronologies of adjacent regions. The German 
and Irish oak chronologies were cross-dated until back into the 3rd millennium BC (Pilcher et al. 
1984), and the German oak chronologies from the Main River, built independently in the Göttingen 
and Hohenheim tree-ring laboratories, cross-date back to 9147 cal BP (Spurk et al. 1998). The North 
American trees that form part of the $^{14}$C calibration curve prior to AD 1320 were cross-dated with 
either the $Sequoiadendron$ master chronology or with a Pacific Northwest Douglas fir chronology 
(Stuiver 1982).

Due to periodic narrow rings caused by cockchafer beetles, some German oak samples were 
excluded from IntCal98. Analysis of these tree rings, with an understanding of the response of trees 
to the cockchafer damage, allowed some of these measurements to be re-instated in the chronology 
(Friedrich et al., this issue).

The 2 parts of the German Preboreal pine chronology (PPC), which were formerly floating, have 
been linked and cross-matched dendrochronologically to the absolutely-dated Holocene oak chro-

nology. Including additional new finds, the south German part of the PPC is prolonged into the
Younger Dryas and now starts at 11,993 cal BP. New pine chronologies from Switzerland and eastern Germany extend the PPC to 12,410 cal BP (Friedrich et al., this issue).

**The Heidelberg Data Set**

$^{14}$C measurements of the wood from the German pine chronology, the Lake Neufchatel chronology, and a new larch series provide terrestrial calibration extending back to 12,410 cal BP. In addition, intervals were measured to fill gaps in the tree-ring $^{14}$C record resulting from previous shifts or cockchafer-damaged sections that could not be re-instated (Friedrich et al., this issue). Additional decadal measurements of German oak have been made for the periods 2570–2800 cal BP and 3440–3640 cal BP. The oak samples were pretreated with acid-alkali-acid (de Vries method), whereas the pines underwent an additional soxhlet extraction to remove resins (Becker and Kromer 1986; Kromer et al. 1986).

**The Belfast Data Set**

Recent $^{14}$C measurements of $\alpha$-cellulose extracted from decadal sections of Belfast Irish oak over the past 1000 yr (Hogg et al. 2002) and for the period 1220–1460 BP (McCormac et al., this issue) have been included. Comparisons of these measurements with the 1986 and 1993 Irish oak data sets (Pearson et al. 1993; Pearson et al. 1986), as well as with Irish oak measurements made in Seattle, resulted in the acceptance of the original 1986 Irish oak data sets over the corrected 1993 data sets (Hogg et al. 2002). Decadal measurements from the 1986 data set have been included separately rather than averaged to bidecadal intervals, where possible. In cases where decadal measurements had been averaged with bidecadal measurements, the original bidecadal data could not be retrieved, so are included as published. In those cases, the decadal data are redundant and are not included. The pre-1993 oak measurements were made on samples processed to holocellulose (Pearson et al. 1986). German oak measurements are used as originally published (Pearson et al. 1993). A further 3 decadal samples of Irish oak from 3450–3470 cal BP were measured to check the earlier Irish oak bidecadal data where they differed considerably from the German oak. These samples were pretreated to $\alpha$-cellulose at Queen's University Belfast and measured at both Queen's University Belfast and the Center for Accelerator Mass Spectrometry, Lawrence Livermore National Laboratory.

**The Waikato Data Set**

Decadal sections of Belfast Irish oak wood from 0–1000 cal BP were pretreated to $\alpha$-cellulose and measured at the Waikato Radiocarbon Laboratory (Hogg et al. 2002).

**The Seattle Data Set**

With the exception of some changes in tree-ring calendar ages in the German pine and oak series (Friedrich et al., this issue) and re-instatement of some cockchafer-damaged sections as described above, the University of Washington Quaternary Isotope Lab data set is the same as detailed in 1998 (Stuiver et al. 1998a). This data set includes wood from Washington, Oregon, California, and Alaska for the period of 0 to 2099 cal BP and from Germany for 1900 to 11,673 cal BP. With the exception of a portion of the Pacific Northwest samples from 35–2095 cal BP which were processed with the de Vries method, all samples were processed to $\alpha$-cellulose.

**The Pretoria and Groningen Data Sets**

$^{14}$C measurements of German oak made at Pretoria and Groningen were published together in 1993 to provide a calibration for short-lived samples (Vogel and van der Plicht 1993). Both laboratories
used a similar acid-base-acid (de Vries method) pretreatment (de Jong et al. 1986; Vogel et al. 1986). The Groningen data set has been revised slightly in the past after discovery of an error in tree-ring counts (de Jong et al. 1989). No changes have been made to these data sets since published in 1993.

The Arizona Data Set

The relation between North American and European wood has been studied using bristlecone pine (BCP) and European oak (German oak and Irish oak), respectively. Discrepancies have become evident over the years, in particular when the German oak was corrected by a dendro-shift of 41 yr towards older ages (Kromer et al. 1996). Attempts were made to resolve the discrepancies by re-measuring BCP samples, measured earlier in Tucson (Linick et al. 1986). The University of Arizona Laboratory of Tree-Ring Research provided dendrochronologically-dated bristlecone pine samples to Heidelberg (wood from around 4700 and 7600 cal BP), Groningen (around 7500 cal BP), Pretoria (around 4900 cal BP), and Seattle (around 7600 cal BP). The replicate measurements have a mean offset of 37 ± 6 14C yr (n = 21) from the Tucson measurements. Applying this shift to the Tucson data results in a close fit to the wiggles of the German oak, which would not occur if there were an error in the dendrochronology of either series. Because of this offset, the IntCal working group has decided not to include the BCP record in IntCal04.

Sample Pretreatment Differences

Tree-ring samples included in IntCal04 have been pretreated either with variations on the de Vries method, which removes sugars, resins, tannins, and part of the lignin, or extracted to holocellulose or α-cellulose as noted above. Because compounds from subsequent years may be deposited in sap-wood, some differences may result in the 14C content of wood pretreated by different methods. Stuiver and Quay (1981) showed that the de Vries method removed about 90% of the later additions in Douglas fir. Since most pre-bomb variations in atmospheric 14C are in the range of 10‰ per year or less, the de Vries method would result in a difference of only a tenth of a mill or <1 14C yr. Oak showed differences in δ13C between samples processed by the de Vries method and that extracted to either holocellulose or α-cellulose, also indicating that not all lignin was removed in the de Vries method (Hoper et al. 1998). The difference in 14C ages resulting from different pretreatment methods is therefore likely to be small, but the difference may contribute to the excess variance observed in interlaboratory comparisons, which are discussed in a subsequent section.

Marine Data Sets, 12.4–26 kyr BP

From 12.4–26 kyr BP, IntCal04 is based on marine data sets with site-specific reservoir corrections. Coral records provide calibration data for the period from 12.4 to 26 cal BP (Bard et al. 1998; Fairbanks et al., forthcoming; Cutler et al., this issue; Burr et al., this issue) with foraminifera records included for the period from 12.4 to 14.7 cal BP (Hughen et al., this issue b). The reservoir corrections for all sites are assumed to have been constant with time with an uncertainty calculated from the overlap with the tree-ring measurements. The reservoir uncertainty is added in quadrature to the measurement uncertainty. The marine data sets and reservoir corrections are discussed in detail in Hughen et al. (this issue b) and Cutler et al. (this issue).

REGIONAL OFFSETS

For the purposes of constructing a Northern Hemisphere calibration data set, the trees used for IntCal04 are assumed to have sampled a well-mixed atmosphere. Latitude-dependent differences in ocean area, surface water ages, and wind speeds are expected to cause 14C offsets of ~1‰ (8 14C yr) for most of the Northern Hemisphere with the exception of the high Arctic (Braziunas et al. 1995).
Measurement of regional offsets in the Northern Hemisphere is difficult because they are near the precision limit for $^{14}\text{C}$. This difficulty is compounded by laboratory offsets (unless samples from each region are measured in the same laboratory and ideally at the same time) and by potential dendrochronological errors. Regional offsets between the primarily mid-latitude sites of the IntCal04 trees have been noted for various time periods, but may not be constant (Stuiver et al. 1998a). These regional offsets were neglected, with the exception of $^{14}\text{C}$ measurements of an Alaskan tree which were normalized to the Pacific Northwest trees by adding the average offset of $14 \pm 3$ yr as reported in Stuiver et al. (1998a). Regional offsets may also result from growing season differences, altitudinal effects, or proximity to areas of intense ocean upwelling, melting permafrost, or volcanic emissions, and are potentially significant for certain localities and time periods (Damon et al. 1996; Deller et al. 2003; Kromer et al. 2001; Manning et al. 2001; Sakamoto et al. 2003). The time-varying nature of such signals, even if they can be detected, precludes an easy correction in the calibration process (Manning et al. 2002). Insufficient data exists at present to construct regional calibration data sets for any extended length of time, with the exception of the Southern Hemisphere. The Southern Hemisphere calibration data set (SHeCal04) and the distinction between Southern and Northern Hemisphere atmospheric boundaries are discussed in more detail in McCormac et al. (this issue).

**SOURCES OF UNCERTAINTY**

Counting statistics do not represent all the uncertainty in $^{14}\text{C}$ ages. Additional sources of uncertainty may be the result of differences in sampling, sample pretreatment, or laboratory operation, which are difficult to quantify for individual samples. The excess variance can be described by comparing the standard deviation in the $^{14}\text{C}$ age difference of replicate samples ($\sigma_2$) with the expected standard deviation based on the counting statistics ($\sigma_1$). One approach to estimating $\sigma_2$ is to assume that it is proportional to $\sigma_1$; it is then convenient to express this relationship through a laboratory error multiplier, $k$, where $k = \sigma_2/\sigma_1$ (Stuiver 1982). These $k$-value estimates should not be used implicitly as a mechanism to determine the added variance in sample handling and analysis relative to pure counting statistics, because not all laboratories base their results only on counting statistics. For instance, the Belfast and Waikato 2002 data sets include laboratory error multipliers of 1.125 and 1.125–1.176, respectively, in the reported errors, and the Arizona NSF accelerator facility and the Center for Accelerator Mass Spectrometry use the larger of the counting statistics uncertainty and the standard deviation in multiple acquisition cycles on the same AMS target to calculate errors on individual samples (Donahue et al. 1997; Guilderson et al. 2003). Regional differences and laboratory practices may also contribute to the variability between laboratories that are not accounted for by laboratory estimates of total error. Nevertheless, for the purpose of constructing a Northern Hemisphere calibration data set, an interlaboratory/inter-regional error multiplier was determined for each of the data sets relative to the Seattle measurements for the same period. The Seattle data set was chosen for this purpose because it includes a large number of replicates and the measurements span most of the period of the tree-ring calibration. An error multiplier of 1.3 was first determined for the Seattle single-year tree-ring data by evaluation of the number of points outside of 2 standard deviations from the mean. Because most of the replicates were re-measured due to some inconsistency with the other data, this error multiplier can be considered an upper limit. This agrees with an upper limit for $k$ for the Seattle data from an earlier comparison between the laboratories that included trees from different regions (Stuiver et al. 1998a). With $k = 1.3$ assigned to the Seattle data, we could then assign the remaining variability to the other data sets. Data for each lab was binned into 10-yr intervals and averaged. Because changes in laboratory procedures and the inclusion of a laboratory error multiplier in recent Belfast results may have affected the interlaboratory error multipliers, we divided the Belfast and Heidelberg data sets, which were measured over the span of several decades, into 2 time
periods for this analysis (Table 1), and also separated the earlier Belfast data set into Irish oak and German oak. There was not a large difference in the calculated \( k \) values between early and recent measurements in the Belfast lab for the Irish oak samples when the previously applied laboratory error multiplier on the more recent data set is considered; however, the early measurements of German oak were more variable than those of Irish oak. The recent Heidelberg data sets had smaller \( k \) values than older measurements. The reason for the early variation is partly due to the fact that these samples were measured to help place a tree in the dendrochronology as it was being built instead of measured consecutively, and also because many of these samples contain only a few tree rings but are being compared to decadal samples. The values of \( k \) in Table 1 were applied to the errors for the tree-ring data sets prior to application of the RWM. Laboratory or data set error multipliers for the marine data sets are discussed in Hughen et al. (this issue a). It will be noted that the Seattle data are generally somewhat younger (positive offset) than the results for the other data sets (Table 1). Unlike IntCal98, no interlaboratory offset corrections were made to the \( ^{14}\text{C} \) ages.

Table 1  Laboratory offsets and error multipliers for the Belfast, Pretoria/Groningen (P/G), Heidelberg, and Waikato laboratories \( ^{14}\text{C} \) measurements of dendro-dated wood compared to Seattle measurements for the same time periods, assuming \( k = 1.3 \) for Seattle (see text). The offset from the Seattle data set equals the weighted mean \( ^{14}\text{C} \) age difference of samples for which the mid-point cal ages fall within the same decade. \( \sigma_1 \) is the predicted average standard deviation in single \( ^{14}\text{C} \) age comparisons (based on quoted laboratory errors); \( \sigma_2 \) is the observed standard deviation in the age and \( n \) is the number of comparisons. The \( \sigma_2/\sigma \) ratio = \( k \). Offset, \( \sigma_1 \), and \( \sigma_2 \) are in \( ^{14}\text{C} \) yr. The \( k \) values marked with an asterisk include previously applied laboratory error multipliers.

<table>
<thead>
<tr>
<th>Laboratory</th>
<th>Offset</th>
<th>( \sigma_1 )</th>
<th>( \sigma_2 )</th>
<th>( k )</th>
<th>( n )</th>
<th>Cal yr interval (cal BP)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Belfast Irish oak 1986</td>
<td>-6 ± 1</td>
<td>21.5</td>
<td>29.4</td>
<td>1.37</td>
<td>407</td>
<td>110–7160</td>
</tr>
<tr>
<td>Belfast German oak 1993</td>
<td>18 ± 2</td>
<td>27.5</td>
<td>48.5</td>
<td>1.76</td>
<td>148</td>
<td>6900–9850</td>
</tr>
<tr>
<td>Belfast Irish oak 2002, 2004</td>
<td>4 ± 2</td>
<td>21.3</td>
<td>27.6</td>
<td>1.21*</td>
<td>124</td>
<td>5–3470</td>
</tr>
<tr>
<td>Pretoria/Groningen 1993</td>
<td>18 ± 2</td>
<td>25.0</td>
<td>27.9</td>
<td>1.12</td>
<td>194</td>
<td>3885–5855</td>
</tr>
<tr>
<td>Heidelberg Hd# &lt;15000</td>
<td>27 ± 2</td>
<td>37.1</td>
<td>50.2</td>
<td>1.35</td>
<td>228</td>
<td>6045–11,655</td>
</tr>
<tr>
<td>Heidelberg Hd# &gt;15000</td>
<td>15 ± 3</td>
<td>31.5</td>
<td>37.0</td>
<td>1.19</td>
<td>86</td>
<td>2575–11,655</td>
</tr>
<tr>
<td>Waikato</td>
<td>10 ± 2</td>
<td>22.0</td>
<td>22.9</td>
<td>1.04*</td>
<td>100</td>
<td>5–995</td>
</tr>
</tbody>
</table>

Uncertainty in single-ring cal ages for dendrochronologically-dated wood is on the order of 1 yr for highly replicated and cross-checked chronologies and is therefore ignored in the analysis. For a \( ^{14}\text{C} \) date on a multiple-year block of wood, there are potentially 2 further effects to be taken into account. Firstly, the \( ^{14}\text{C} \) measurement in such a case is an average over a period in which the true curve will have changed by an unknown amount; the uncertainty that this adds is taken into account in the RWM, which explicitly allows for year-on-year changes in the underlying curve (see Buck and Blackwell, this issue), unlike in IntCal98 where it was essentially ignored. Secondly, this average may be based on an unequal amount of carbon from each ring; if the actual proportions are unknown, this introduces an additional small uncertainty. A worst-case offset of 2 yr in calendar age was determined from the annual tree-ring data (Stuiver et al. 1998) in terms of moving decade age compared to IntCal98 (Manning et al. 2001). A comparison of annual to decadal tree-ring \( ^{14}\text{C} \) measurements of a single Douglas fir tree from 440–250 cal BP (see supplemental data) gives a mean offset of 5.5 ± 4.4 \( ^{14}\text{C} \) yr. The larger uncertainty due to loss of wood from saw cuts in some early
sample preparations (M Friedrich, personal communication) is more critical and hard to define because wood often does not have exactly vertical ring divisions, so sawing could include or exclude one or even several rings for part of the sample under the initial surface. This is especially likely for narrow rings or for smaller tree sections where curvature comes into play. Dissecting of tree rings can be difficult depending on the species and size of rings, but with care only a small amount of wood from the current ring is lost or additional wood from an adjacent ring included. At present, these effects are ignored for simplicity, although the RWM could readily incorporate such information if it was available.

CALIBRATION CURVE CONSTRUCTION

IntCal98 was constructed by taking a weighted average of all data within a 10-yr window and assigning the mid-point of the decade as the cal age. Bidecadal tree-ring Δ14C measurements were treated as if there had been 2 decadal measurements, and sub-decadal measurements were treated as if they were decadal. The uncertainty in the calendar age of varved sediments and U/Th-dated coral was not considered. For IntCal04, a coherent statistical method was used (Buck and Blackwell, this issue) based on a stochastic model for the underlying 14C calibration curve, taking into account the calendar age span (e.g. number of tree rings or varves) and calendar age uncertainty of the samples as well as 14C age uncertainty. The estimation allows for the covariance between observations due to the structure of the model and to dependence between the calendar age uncertainty of related observations (e.g. those with varve-counting or wiggle-matching errors in common). The model for the curve takes the form of a random walk, with the changes in the curve from one year to the next represented by a Gaussian distribution with a mean of β and variance per year of r². Because there are usually many observations relevant to the estimated 14C age for a given calendar year, the calculation is carried out within a window of at least 100 observations. The value for the parameter r was derived from sensitivity tests on single-year tree-ring measurements of the past 500 yr (Stuiver et al. 1998b). The single-year tree-ring measurements were initially selected for the model parameterization in order to retain as much signal as possible, but essentially the same value for r was obtained using decadal measurements in selected portions of the tree-ring data set. The parameter selection details are given in Buck and Blackwell (this issue). The IntCal04 14C calibration curve is generated by the model at intervals of 5 yr for the range 0–12.4 cal kyr BP, 10 yr for 12.4–15 cal kyr BP, and 20 yr for 15–26 cal kyr BP.

DISCUSSION

For comparison purposes, we calculated a weighted average of all updated tree-ring 14C measurements within a 10-yr window. This 10-yr binned data set is equivalent to the procedure used to construct IntCal98. The RWM assumes that there is equal probability of a rise or fall in atmospheric 14C levels around the mean value of β. Because 14C production increases may be more rapid than the subsequent redistribution in the global carbon reservoirs, we tested the validity of this assumption by calculating the distribution of shifts from one decade to the next. These shifts are normally distributed in both the binned tree-ring data set and IntCal04 (Figure 2), indicating that the assumption is valid.

To evaluate the effect of the RWM on the periodic nature of the 14C data, we examined the spectral properties of the Δ14C calculated from the IntCal04 calibration curve from 2400–100 BP, so that no detrending was necessary. Both data sets contained significant spectral power at about 90 yr and 200 yr with approximately the same amplitude using a fast fourier transform with confidence limits (Mitchell et al. 1966). The 205-yr cycle was phase-shifted by –0.3 degrees and the 88-yr cycle by +10.8 degrees in IntCal04 relative to the binned tree-ring data set, which indicates that geophysical
variations on decadal and centennial scales are preserved in IntCal04. Attenuation of the signal and a larger phase shift for higher-frequency components is possible, so caution is in order when comparing cycles with other records.

Wiggle-matching to the IntCal04 calibration data set may require some adjustment in the methods in use, particularly with shorter series. The fit for shorter series may also be significantly affected by the smoother curve of IntCal04 relative to IntCal98. The more general issue here is in the methodology used for fitting. The usual underlying assumption is that the points on the calibration curve can be treated as independent observations, whereas, in fact, they are syntheses of possibly overlapping data sets. In principle, this will be a problem for any analysis where the terms in the likelihood function are combined as if this assumption of independence were true. In practice, the uncertainties in the measurements are usually much higher than those in the curve, so this is not of any practical significance. In the case of wiggle-matching of tree-ring sequences, the method is sometimes being pushed to the limits in all respects. In these cases, it may be necessary to incorporate the fitted data and the original calibration data in a full Bayesian analysis (see Buck and Blackwell, this volume); such an approach could of course be used in all cases.

CONCLUSIONS

Because the RWM accounts for scatter around the “true calibration curve,” the IntCal04 curve is somewhat smoother than the tree-ring-derived portion of IntCal98 (0–11,440 cal BP), except where there are measurements on 1–3-yr pieces of wood (Figure 3a). The addition of new data in certain time periods also results in tighter constraints on the curve in places (Figure 3b). The most significant changes (Figure 3c–e) are of course due to the extension of the German dendrochronology from 11.4 to 12.4 cal kyr BP (Friedrich et al., this issue), the new high-resolution Cariaco Basin data set from
10.5 to 14.7 cal kyr BP (Hughen et al., this issue b; Hughen et al. 2000), and extensive new coral data sets from 12.4 to 26 cal kyr BP (Fairbanks et al., forthcoming; Cutler et al., this issue; Burr et al., this issue). The resulting difference in calibrated age ranges may be substantial in these cases—e.g., a hypothetical $^{14}$C age of $11,000 \pm 100$ $^{14}$C BP yields 95% cal age ranges of cal BP 12,650–12,730 and 12,810–13,190 with IntCal98, and cal BP 12,820–13,100 with IntCal04 (Figure 4).

Figure 3a IntCal98 (gray lines) and IntCal04 (black lines) 1-standard deviation error envelopes for 0–500 cal BP.

Figure 3a IntCal98 (gray lines) and IntCal04 (black lines) 1-standard deviation error envelopes for 0–500 cal BP.
Figure 3b IntCal98 (gray lines) and IntCal04 (black lines) 1-standard deviation error envelopes for 2500–3000 cal BP.
Figure 3c IntCal98 (gray lines) and IntCal04 (black lines) 1-standard deviation error envelopes for 11,500–12,500 cal BP.
Figure 3d  IntCal98 (gray lines) and IntCal04 (black lines) 1-standard deviation error envelopes for 12,500–14,500 cal BP.
Figure 3e IntCal98 (gray lines) and IntCal04 (black lines) 1-standard deviation error envelopes for 16,000–26,000 cal BP.

Figure 3e IntCal98 (gray lines) and IntCal04 (black lines) 1-standard deviation error envelopes for 16,000–26,000 cal BP.
The entire IntCal04 data set and calibration curve is plotted in 2000-yr sections for use in estimating calibrated age ranges (Appendix, Figure A1–A13). Note that there are still gaps of greater than 700 yr in the calibration data at 16–17 cal kyr BP, 20.6–21.6 cal kyr BP, and 23.8–24.6 cal kyr BP. 14C age calibrations in these sections will therefore be less precise; this is reflected in the increased uncertainty on the curve in these regions.

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**Figure 4** Calibrated probability distributions for a hypothetical 14C age of 11,000 ± 100 14C BP calibrated with IntCal98 (light gray) and IntCal04 (dark gray) showing 95% cal age ranges.
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Pearson GW, Becker B, Qua F. 1993. High-precision 14C measurement of German and Irish oaks to show the natural 14C variations from 7890 to 5000 BC. *Radiocarbon* 35(1):93–104.


Figure A1 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}\text{C}$ increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A2 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A3 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A4 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A5 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A6 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A7. The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A8: The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A9 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A10 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A11 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A12 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.
Figure A13 The IntCal04 terrestrial calibration curve (1-standard deviation envelope) and data with 1-standard deviation error bars in $^{14}$C increased by the laboratory error multipliers described in the text. The uncertainty in the calendar ages is not shown, but is taken into account in the random walk model.